## Estimation of conditional mutual information and its application as a measure of conditional dependence

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Conditional independence is an important concept in many different aspects of engineering such as detecting causal relationship between two time series and extracting informative features in a classification or regression setting [2, 3, 5]. Besides several other measures, conditional mutual information (CMI) has been suggested as a measure of conditional dependence [4]. Inspired by the use of kernel based methods as a function approximator, we propose a new approach to estimate CMI. We briefly describe the proposed work below.

Given random variables (X, Y, Z), CMI between X and Y given Z is defined as

$$I(X;Y|Z) = \mathbb{E}_Z[I(X;Y)|Z] = \int f_Z(z) \left[ \iint f_{XY|Z}(x,y|z) \log \frac{f_{XY|Z}(x,y|z)}{f_{X|Z}(x|z)f_{Y|Z}(y|z)} \mathrm{d}x\mathrm{d}y \right] \mathrm{d}z.$$
(1)

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It can be easily seen that conditional mutual information can be expressed in the following form,

$$I(X;Y|Z) = \mathbb{E}_{XYZ} \left[ \log \left( \frac{f_{XYZ}(X,Y,X)}{f_X(X)f_Y(Y)f_Z(Z)} \frac{f_X(X)f_Z(Z)}{f_{XZ}(X,Z)} \frac{f_Y(Y)f_Z(Z)}{f_{YZ}(Y,Z)} \right) \right]$$

Let us define

$$h(x, y, z) = \frac{f_{XYZ}(x, y, z)}{f_X(x)f_Y(y)f_Z(z)}.$$

Then

$$f_{XYZ}(x, y, z) = h(x, y, z)f_X(x)f_Y(y)f_Z(z)$$

Integrating both sides with respect to x we get,

$$f_{YZ}(y,z) = f_Y(y)f_Z(z)\mathbb{E}_X[h(X,y,z)] \Rightarrow \frac{f_{YZ}(y,z)}{f_Y(y)f_Z(z)} = \mathbb{E}_X[h(X,y,z)].$$

Similarly,

$$\frac{f_{XZ}(x,z)}{f_X(x)f_Z(z)} = \mathbb{E}_Y[h(x,Y,z)].$$

Thus,

$$I(X;Y|Z) = \mathbb{E}_{XYZ} \left[ \log h(X,Y,Z) - \log \mathbb{E}_{X'} [h(X',Y,Z)] - \log \mathbb{E}_{Y'} [h(X,Y',Z)] \right]$$
(2)

where X' and Y' are independent copies of X and Y respectively. Therefore, an estimator of h(x, y, z) directly leads to an estimator of CMI.

Inspired by the kernel based methods, we replace h by  $\hat{h}(x, y, z) = \sum_{i=1}^{n} \alpha_i \kappa_1(x, x_i) \kappa_2(y, y_i) \kappa_3(z, z_i)$ where  $\kappa_1, \kappa_2$  and  $\kappa_3$  are positive definite kernels and minimize the following cost function solve for the coefficients,

$$J = \int \left(\frac{f_{XYZ}(x, y, z)}{f_X(x)f_Y(y)f_Z(z)} - \hat{h}(x, y, z)\right)^2 f_X(x)f_Y(y)f_Z(z)dxdydz + \lambda ||\alpha||_2^2$$

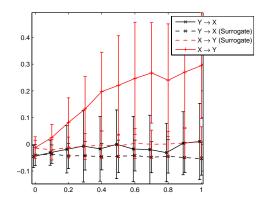


Fig. 1: Surrogate data test of causality.

where  $\alpha$  is a vector containing the coefficients  $\alpha_i$ 's. Expanding the first term in the cost function it can be shown that

- 1. the actual density functions only appears as expectations i.e. we do not need to know them explicitly and,
- 2. the cost function is quadratic in terms of  $\alpha$  and therefore, can be easily solved.

Replacing  $\hat{h}$  in eq. (2), we get an estimator of CMI. Next, we briefly describe an experiment of detecting causality.

Consider two time series  $\{X_t\}$  and  $\{Y_t\}$  as follows,

$$\begin{aligned} x(t) &= 3.4x(t-1)(1-x^2(t-1))e^{-x^2(t-1)} + 0.8x(t-2) + \epsilon_1 \\ y(t) &= 3.4y(t-1)(1-y^2(t-1))e^{-y^2(t-1)} + 0.5y(t-2) + cx^2(t-2) + \epsilon_2 \end{aligned}$$

where  $\epsilon_1, \epsilon_2 \sim \mathcal{N}(0, 1)$ . Similar experiment without the noise term has been described in [1]. In this particular case  $\{X_t\}$  causes  $\{Y_t\}$  but not the other way around. Therefore,  $X_t \perp [Y_{t-1}, Y_{t-2}]|[X_{t-1}, X_{t-2}]$  but  $Y_t \not\perp [X_{t-1}, X_{t-2}]|[Y_{t-1}, Y_{t-2}]$ . Figure 1 shows the surrogate data test with 100 samples. It can be clearly seen that the test supports our expectations.

## References

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